Signal Processing for CBR defence

Branko Ristic¹

DST group, Australia



Australian Government Department of Defence Defence Science and Technology Organisation

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¹With contributions from A. Skvortsov, A. Gunatilaka, R. Gailis, M. Roberts, A. Walker, M. Morelande

• CBR defence: protective measures against a

- chemical (C)
- biological (B)
- radiological (R)

attack.

- CBR defence covers a fairly broad spectrum of measures; multi-disciplinary scientific effort
- This talk will focus on signal processing techniques applied to recent and ongoing projects within the Land Division of DST group

Localisation of a source of biochemical agent release

- Q Autonomous search techniques for CBR sources
- **③** Image reconstruction for a standoff gamma radiation detector
- Forecasting of an epidemic outbreak
- Section of biological systems with competition

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Localisation of a biochemical source: Context

- Deliberate or accidental release of toxic material (gas, aerosols) into the atmosphere
- Problem: Given the readings supplied by a network of spatially distributed biochemical sensors, estimate the location of the source
- Technical difficulties:
 - source strength unknown;
 - characteristics of sensor measurements not well understood (plethora of air dispersion models; likelihood function only a guess)





Localisation of a biochemical source: Problem formulation

- Parametric approach parameter vector $\boldsymbol{\theta}$ includes
 - Source location (x_0, y_0, z_0) , strength Q_0 (possibly size)
 - Enviro/meteo parameters
 e.g. wind speed, direction, canopy characteristics
- Concentration measurements at spatially distributed sensor locations $z_i \geq 0, i = 1, \dots, S$
 - Mean concentration at sensor $i: \mathbb{E}\{z_i\} = h_i(\theta)$... based on the adopted dispersion model
 - Random fluctuations (capture both the dispersion modelling errors and measurement noise)
 - Possibly quantised
- Bayesian estimation framework

$$\underbrace{p(\boldsymbol{\theta}|z_1,\ldots,z_S)}_{\text{posterior PDF}} \propto \underbrace{\ell(z_1,\ldots,z_S|\boldsymbol{\theta})}_{\text{likelihood function}} \cdot \underbrace{\pi(\boldsymbol{\theta})}_{\text{prior PDF}}$$

Due to turbulence:

- meandering of the plume
- intermittent measurements



(diffusion)



(diffusion + advection)

Two categories of transport of substances in the environment

- Advection transport with the mean fluid flow
- Diffusion transport through the action of random motions

- Previous approaches:
 - Adopt a dispersion model (which determines $h_i(\boldsymbol{\theta})$)
 - Adopt a certain model of random measurement fluctuations (Gaussian, log-normal)
 - Use Markov chain Monte Carlo (MCMC) to approximate the posterior $p(\pmb{\theta}|z_1,\ldots,z_S)$
- Proposed framework²:
 - Adopt several candidate dispersion models: $\mathcal{M} = \{1, \dots, M\}$
 - Likelihood-free estimation
 - Use approximate Bayesian computation (ABC) to simultaneously estimate $p(m|z_1,\ldots,z_S)$ and $p(\theta_m|m,z_1,\ldots,z_S)$

² Ref: B. Ristic et al. 'Bayesian likelihood-free localisation of a biochemical source using multiple dispersion models', Signal Processing, 2015

Multiple-model ABC rejection sampler

- $\pi_{\mathcal{M}}(m)$... prior PDF over dispersion models $m = 1, 2, \dots, M$
- $\pi_{\pmb{\theta}_m}(\pmb{\theta}_m)$... prior PDF over the parameter space for model m

1: Input:
$$\mathbf{z} = [z_1, \dots, z_S]^{\mathsf{T}}$$
; ϵ ; N
2: Initialise: $\mathcal{X}_1 = \dots = \mathcal{X}_M = \emptyset$
3: repeat
4: Draw m^* from $\pi_{\mathcal{M}}(m)$
5: Draw θ_m^* from $\pi_{\theta_m^*}(\theta_m)$
6: Simulate measurement \mathbf{z}^* using model m^* and parameter θ_m^*
7: Compute distance $d^* = D(\mathbf{z}, \mathbf{z}^*)$
8: if $d^* \leq \epsilon$ then
9: $\mathcal{X}_{m^*} = \mathcal{X}_{m^*} \cup \{\theta_m^*\}$
10: end if
11: until $\sum_{m=1}^M |\mathcal{X}_m| = N$
12: Output: $\mathcal{X}_1, \dots, \mathcal{X}_M$

 $p(m|\mathbf{z}) \approx \frac{|\mathcal{X}_m|}{N}; \quad p(\boldsymbol{\theta}_m|m,\mathbf{z}) \text{ approximated by the sample } \mathcal{X}_m$

- Rejection sampling is inefficient (very low acceptance rate)
- We developed a more efficient scheme: rejection sampling is applied using a sequence of monotonically decreasing tolerance levels
 ε₁ > ε₂ > · · · > ε_T > 0;
- The subsequent approximations gradually approach the true posterior
- Similar to the SMC-ABC sampler (Toni et al., 2009), but no need to specify in advance tolerances $\epsilon_1, \ldots, \epsilon_T$; instead, iteration t computes the tolerance for the next iteration, ϵ_{t+1}

• M = 3 models considered:

- m=1: Gaussian plume with a linear spread, dim $({m heta}_1)=7$
- m=2: Gaussian plume with a nonlinear spread, dim $(\boldsymbol{\theta}_2)=9$
- m = 3: Stretch exponential model, dim $(\theta_3) = 10$

• Gaussian plume model

- Gaussian distribution of the plume in the vertical and horizontal directions.
- By convention, x-axis coincides with the average wind velocity vector
- Mean concentration at sensor \boldsymbol{i}

$$C_{i} = h_{i}(\boldsymbol{\theta}) = \frac{Q_{0}}{2\pi\sigma_{y_{i}}\sigma_{z_{i}}U} e^{-\frac{(y_{i}-y_{0})^{2}}{2\sigma_{y_{i}}^{2}}} \left[e^{-\frac{(z_{i}-z_{0})^{2}}{2\sigma_{z_{i}}^{2}}} + e^{-\frac{(z_{i}+z_{0})^{2}}{2\sigma_{z_{i}}^{2}}}\right]$$

with $\sigma_{y_i}=rac{\sigma_v}{U}(x_i-x_0)$, $\sigma_{z_i}=rac{\sigma_w}{U}(x_i-x_0)$ for m=1 (point source)

Experimental dataset for algorithm evaluation

- Collected using a recirculating water channel (COANDA R&D Corp), 10m × 1.5m × 0.9m
- The source: constant release of fluorescent dye from a narrow tube
- Concentration measurements collected at S=48 downstream positions using laser induced fluorescence (LIF); averaged over 100 sec
- The floor of the w/c covered with a mesh to give surface roughness
- Two scenarios considered:
 - without obstacles (open terrain);
 - with obstacles (urban environment)



Experimental dataset: Results (1)



Top-down view of the experimental setup





Localisation at iterations t = 1, 4, 7, 10

Experimental dataset: Results (2)



Top-down view of the experimental setup





Localisation at iterations t = 3, 6, 9, 12

Localisation of a biochemical source: CR bounds (1)

Assuming:

- Gaussian plume dispersion model
- Additive white Gaussian measurement noise

Derived the theoretical Cramer-Rao lower bounds for estimation 3 of heta

• analog measurements: $z_i = h_i(\boldsymbol{\theta}) + w_i$,

• binary measurements:
$$b_i = \begin{cases} 1, & z_i > \tau \\ 0, & z_i \leq \tau \end{cases}$$
, $i = 1, \dots, S$

analog

Localisation information

binary



³Ref: B. Ristic et al., "Achievable accuracy in Gaussian plume parameter estimation ...", Information Fusion, 2015.

Localisation of a biochemical source: CR bounds (2)



Problem:

- Identical spatially distributed binary sensors
- Binary threshold au is unknown (vague prior)
- Semiurban environment

Question: Can we localise the source (estimate θ) in this case (since the source strength Q_0 is also unknown)?

Answer: Yes, replace Q_0 (in θ) with Q_0/τ .

Localisation using experimental binary data (1)

- State vector: $\boldsymbol{\theta} = [x_0 \quad y_0 \quad Q_0/\tau \quad U]^{\intercal}$
- Dispersion model: the count of particle encounters in 2D (Vergassola et al, Nature, 2007)
- Likelihood function for the measurement vector $\mathbf{b} = [b_1, \dots, b_S]^{\mathsf{T}}$ derived assuming that z_i is Poisson distributed

Prior distributions:





Localisation using experimental binary data (2)



- Estimation carried out using the importance sampling technique
- Similar results obtained for three experimental datasets

⁴ Ref: B. Ristic et al. 'Source Localisation in turbulent flow using a binary sensor network', *IEEE SPL* (in review), 2015

Localisation of a source of biochemical agent release

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- Consider sensors that can be controlled, e.g.
 - where to go
 - how long to stay
- Search is a repetitive cycle of sensing, estimation (localisation), and sensor control
- Decisions (controls, actions) are made sequentially and autonomously (AI) in the presence of uncertainty, using only past measurements and past actions.
- Goal: minimise the search time via on-line control of individual sensors (need to define in advance the termination criterion)

- Search strategies for finding a source of an emission (gas, particles, smell, energy) based on sparse cues are of great importance
 - National security (e.g. release of hazardous substances)
 - Recovery & rescue emissions (e.g. MH370)
 - Understanding nature (foraging behaviour of animals)
- Very popular subject; huge amount of literature (top journals); various scientific disciplines (e.g. biology, physics, operations research, robotics)
- Categories:
 - Deterministic (systematic search) vs random search
 - Random search models: diffusion, Lévy flight, intermittent search (exploration vs exploitation)
 - Gradient search (chemotaxis) vs information driven search (infotaxis)

Sequential decisions under uncertainty

"Partially observed Markov decision problem" formulation

The elements of POMDP:

- A set of admissible actions (controls) at discrete-time k: \mathbb{U}_k ;
- The (information) state of the system at time k: \mathcal{I}_k
- A reward function associated with each action $\mathbf{u} \in \mathbb{U}_k$: $\mathcal{R}(\mathbf{u}, \mathcal{I}_k)$

The best action should be based on the reward computed a few steps ahead.

One step ahead optimal sensor control:

$$\mathbf{u}_k = \arg \max_{\mathbf{v} \in \mathbb{U}_k} [\mathcal{R}(\mathbf{v}, \mathcal{I}_k)]$$

Bayesian - Information Theoretic Approach

• The (information) state of the system is the *posterior PDF*:

$$\mathcal{I}_k = p(\boldsymbol{\theta} | \mathbf{z}_{1:k}, \mathbf{u}_{0:k-1}), \qquad (\mathbf{z}_{1:k} \equiv \mathbf{z}_1, \mathbf{z}_2, \dots, \mathbf{z}_k)$$

- Sequential Bayesian estimation; particle filter implementation
- Reward function: an information theoretic measure, e.g.
 - Entropy difference: $\begin{aligned} \mathcal{R} &= \mathbb{E}\{H_{k+1}(\mathbf{u},\mathcal{I}_{k+1})\} H_k(\mathcal{I}_k) \\ & \text{where } H_k(\mathcal{I}_k) = -\int p(\boldsymbol{\theta}|\mathbf{z}_{1:k},\mathbf{u}_{0:k-1}) \ln p(\boldsymbol{\theta}|\mathbf{z}_{1:k},\mathbf{u}_{0:k-1}) d\boldsymbol{\theta} \end{aligned}$
 - Rényi divergence

$$\mathcal{R} = \mathbb{E}\left\{\frac{1}{\alpha - 1}\log\int [p(\boldsymbol{\theta}|\mathbf{z}_{1:k+1}, \mathbf{u}_{0:k})]^{\alpha}[p(\boldsymbol{\theta}|\mathbf{z}_{1:k}, \mathbf{u}_{0:k-1})]^{1 - \alpha}d\boldsymbol{\theta}\right\}$$

($\alpha=1$ \Rightarrow Kullback-Liebler div.; $\alpha=0.5$ \Rightarrow Bhattacharyya dist.)

- Turbulent diffusion
- Real (COANDA) data



Autonomous search demo: Explanation

- Likelihood function: Poisson distribution
- Dispersion model: the count of particle encounters in 2D (Vergassola et al., Nature, 2007); the mean of the Poisson distribution at sensor i

$$C_i = h_i(\boldsymbol{\theta}) = \frac{Q_0}{\ln\left(\frac{\lambda}{a}\right)} \exp\left[\frac{(x_0 - x_i)U}{2D}\right] \cdot K_0\left(\frac{\sqrt{(x_i - x_0)^2 + (y_i - y_0)^2}}{\lambda}\right)$$

where $\lambda=\sqrt{(DT_o)/(1+\frac{U^2T_o}{4D})}$, a - size of the sensor; D - diffusivity; T_0 - mean lifetime of a particle; U - mean wind speed

• The reward function: Entropy difference

Autonomous search: study⁵

- All information theoretic based reward functions perform similarly.
- An important factor in search performance is the ratio $\rho = \frac{A_{search}}{A_{annu}}$

The smaller ρ , the more efficient information driven search strategy (infotaxis) is compared to the systemat search (and vice versa)



⁵B. Ristic et al., A study of cognitive strategies for an autonomous search using sparse cues, *Information Fusion*, 2015

Let the reward function be the entropy difference.

Postulation: Dynamic model of entropy during the search is a Gauss-Markov random walk with a constant (negative) drift.



Brownian motion (as a continuous-time limiting case of a Gauss-Markov random walk) has the following property:

Distribution of the "first passage time"

The time which Brownian motion $X_t = \nu t + \sigma W_t$, with a drift $\nu < 0$, takes to reach a certain (negative) level, is an inverse Gaussian RV

The PDF of search time (2)



QQ plots of 4 samples of search time (varying U and Q_0)



Autonomous search for radiological point sources⁶

- Sensor: low-cost, non-directional (Geiger counter)
- Unknown number of sources R ≥ 0; unknown intensities
- Control vectors: where to go, how long to stay
- Propagation model: inverse distance squared

$$C_{i} = h_{i}(\boldsymbol{\theta}) = \sum_{r=1}^{R} \frac{Q_{0}^{(r)}}{\left[x_{i} - x_{0}^{(r)}\right]^{2} + \left[y_{i} - y_{0}^{(r)}\right]^{2}}$$

 Likelihood function: Poisson, i.e. z_i ~ P(h_i(θ) + β)



Top down view of the data experimental set-up

⁶B. Ristic et al., Information driven search for point sources of gamma radiation, *Signal Processing*, 2010

Autonomous search for radiological point sources (cont'd)



Autonomous search in an unknown structured environment

- Context⁷:
 - A diffusive source of toxic substance (gas, particles)
 - Search domain contains obstacles
 - GPS denied environment; the map is unknown
 - The initial position and the search domain boundary are given
- Goal: Estimate the source location and the path to it!
- Assumption: The searcher can sense

- the concentration (of toxic substance)
- 2 the presence / absence of obstacles



⁷B. Ristic et al., Autonomous search for a diffusive source in an unknown structured environment, *Entropy*, 2014

Discretisation of the search area



Complete lattice (known a priori) Incomplete lattice (unknown) (obstacles are missing links, $\sim 35\%$) (percolation threshold)

х

5

10

- The correct likelihood function depends on the map because the map is unknown, the searcher uses an approximation derived using the conformal mapping technique
- Execution of motion controls prone to error (with prob $p_e \ll 1$)
- The state vector $\boldsymbol{\theta}$ includes:
 - **O** Coordinates (x_0, y_0) , intensity of the source Q_0
 - The map (i.e. existence of all links in the lattice)
 - In the searcher position on the map
- Admissible actions: $\mathcal{U}_k = \{\cdot, \rightarrow, \leftarrow, \uparrow, \downarrow\}$
- Reward function: Bhattacharyya distance (over x_0, y_0, Q_0 only)

Search - unknown structured environment: demo (2)





Examples on a bigger scale





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Gamma radiation image reconstruction⁸

- DSTO recently built a prototype of a standoff imaging gamma radiation detector
- Purpose: determine exact locations of radiological sources within the field of view
- Rotational modulation collimation
 - Two attenuating masks separated by a known distance
 - Co-rotating on a cylinder in front of three gamma ray detectors
- Problem: Reconstruct the image from its RMC projection





⁸Ristic, Roberts, "A parametric Bayesian RMC gamma-ray image reconstruction", ICASSP 2015

• Non-parametric formulation

- Emission tomography (reconstruction of medical images)
- Standard algorithms: EM, MAP

Measurements:
$$y_i \sim \mathcal{P}_{y_i}\left(\sum_{j=1}^N A_{ij}\lambda_j\right)$$
, (rot. angles $i = 1, \dots, M$)

• Goal: reconstruct image $\lambda = [\lambda_1, \dots, \lambda_N]^{\intercal}$ using $\mathbf{y} = [y_1, \dots, y_M]^{\intercal}$

• Parametric formulation

- Motivation: limited projections (8); simpler images (9)
- Image: a weighted sum of Gaussian radial basis functions:

$$\lambda_j(\boldsymbol{\theta}) = \sum_{k=1}^Q \lambda_{jk}(\boldsymbol{\theta}), \text{ where } \lambda_{jk}(\boldsymbol{\theta}) = w_k e^{-\frac{(x_j - \bar{x}_k)^2 + (y_j - \bar{y}_k)^2}{\sigma_k^2}}$$

• Goal: reconstruct the image by estimating the 4Q-dimensional parameter vector $\boldsymbol{\theta} = [\mathbf{w}^{\mathsf{T}} \ \bar{\mathbf{x}}^{\mathsf{T}} \ \bar{\mathbf{y}}^{\mathsf{T}} \ \boldsymbol{\sigma}^{\mathsf{T}}]^{\mathsf{T}}$

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Bayesian parametric RMC image reconstruction

- Bayesian parameter estimation framework: $\pi(\theta|\mathbf{y}) \propto p(\mathbf{y}|\theta) \pi_0(\theta)$
- Implementation: importance sampling with progressive correction and Metropolis-Hastings step
- The proposed method works much better for extended (non-point) sources
- Work in progress: model selection for unknown Q; fast estimation(to deal with large Q, e.g. non-homogeneous background radiation)







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Dynamics of infectious disease

- Epidemic:
 - a complex dynamic stochastic system
 - on the macroscopic level: compartmental models of disease transmission
- Example: SEIR (compartmental) model (measles, pox, influenza, ...)



(ignoring birth, migration, death by natural causes)

$$\begin{split} \dot{S} &= -\beta IS/N \\ \dot{E} &= \beta IS/N - \alpha E \\ \dot{I} &= \alpha E - \gamma I \\ R &= N - (S + E + I) \end{split}$$



Sequential Epidemic State Estimation (and Prediction)

- Nonlinear (stochastic) filtering framework
- Stochastic SEIR model: Gillespie algorithm, approximations

$$\begin{split} S_{k} &= S_{k-1} - \nu_{k}, \qquad \nu_{k} \sim \mathcal{P}(\beta I_{k-1} S_{k-1} \Delta / N) \underbrace{\mathbb{P}_{2000}^{2500}}_{I_{k} = I_{k-1} + \nu_{k} - \mu_{k}, \qquad \mu_{k} \sim \mathcal{B}(E_{k-1}, e^{-\Delta \alpha}) \\ I_{k} &= I_{k-1} + \mu_{k} - \eta_{k}, \qquad \eta_{k} \sim \mathcal{B}(I_{k-1}, e^{-\Delta \gamma}) \underbrace{\mathbb{P}_{2000}^{2500}}_{500} \underbrace{\mathbb{P}_{2000}^{1000}}_{0 \to 0} \underbrace{\mathbb{P}_{2000}^{1000}}_{0 \to 0} \underbrace{\mathbb{P}_{2000}^{1000}}_{0 \to 0} \underbrace{\mathbb{P}_{2000}^{1000}}_{0 \to 0} \underbrace{\mathbb{P}_{2000}^{1000}}_{1 \to 0} \underbrace{\mathbb{P}_{2000}^{1000}}_{0 \to 0} \underbrace{\mathbb{P}_{2000}^{100$$

- Measurements
 - Cumulative number of infected cases (e.g. Ebola)
 - Syndromic surveillance (e.g. flu number of Google searches)
- Implementation: particle filter⁹

⁹Skvortsov, Ristic, "Monitoring and prediction of an epidemic outbreak", Math. Biosciences, 2012

Application: Ebola virus epidemic in west Africa

- WHO data (available online)
- Likelihood function unknown: Neg-Binomial vs Poisson



¹⁰Ristic, Dawson, "Forecasting an Epidemic Outbreak: Application to Ebola Cases", *Signal Processing* (in review), 2015

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- Extinction of biological systems with competition

Extinction and survival in a competitive world

- Central themes of population biology
- Relevant in a broader context of complex stochastic dynamic systems (e.g. virus evolution, stock market trading)
- Mathematical model: single-prey multiple-predator Lotka-Volterra (LV-1n) system
- prey population (food for predators): \boldsymbol{x}
- competing predators: y_1, \ldots, y_n

$$\dot{x} = \alpha x \left(1 - \sum_{i=1}^{n} \beta_i y_i \right)$$

$$\dot{y}_i = \beta_i x y_i - \gamma_i y_i$$

 $\alpha, \beta_1, \gamma_1, \ldots, \beta_n, \gamma_n > 0$

Deterministic model - no extinction!



PDF of extinction time for a stochastic LV-1n

2n+1 biochemical reactions:

$$\begin{array}{cccc} X & \stackrel{\alpha}{\to} & 2X \\ X + Y_i & \stackrel{\beta_i}{\to} & 2Y_i & i = 1, \dots, n \\ Y_i & \stackrel{\gamma_i}{\to} & \emptyset & i = 1 & n \end{array}$$

Master eq. \Rightarrow exact simulation algorithm (Gillespie)

- We are after the PDF of extinction time T
- Simplification: competing predators collapse into one aggregated predator y (parameters $\beta,\,\gamma)$
- Analytic expression for the PDF of $\tilde{T}=T/\tau$ is then known (Kamenev, Parker, 2009)

$$p(\tilde{T}) = \frac{a}{\sqrt{\pi \tilde{T}^3}} \exp\left[-(\tilde{T} - a)^2/\tilde{T}\right]$$

where: τ is a (known) function of x(0), y(0), α , β and γ ; $a \approx 0.5$



PDF of extinction time: example

Suppose we observe occasionally, over a period of time, the prey count of an LV-1n system¹¹(N.B. n can be unknown!)

• Assuming:

- an aggregated stochastic LV-11 model
- Poisson likelihood of prey count measurements

we can estimate $x(0),\,y(0),\,\alpha,\,\beta$ and γ (e.g. using the pMCMC algorithm), find τ and the PDF of T



¹¹Ristic, Skvortsov, "Predicting extinction of a biological system with competition", Chapter 25 in Emerging Trends in Computational Biology, Bioinformatics, and Systems Biology, 2015

- Review signal processing techniques applied to a few recent and ongoing CBR defence projects in Land Division of DSTO
- Future work:
 - Practical: build autonomous search robots
 - Theoretical: distributed multi-platform search (estimation, control), intermittent search strategies, parameter estimation of large stochastic reaction networks (system biology: signaling pathways)

Questions ???